

On the new unit Lindley distribution

Subrata Chakraborty

Department of Statistics, Dibrugarh University, Assam, India

ARTICLE HISTORY

Compiled June 22, 2021

Reeived 27 February 2021; Accepted 04 June 2021

ABSTRACT

In this article it is shown that the new unit Lindley distribution of Mazuchelli et al. [2] is in fact complementary of the original unit Lindley distribution of Mazuchelli et al. [1] and as such almost all properties of the former directly follows from those of the later quite easily. Some illustrative derivation of properties of new unit Lindley distribution from those of the unit Lindley distribution is also presented.

KEYWORDS

Complementary random variable; Unit-Lindley distribution ; Exponential family

1. Introduction

Recently, [2] introduced a new unit Lindley (NUL) distribution and studied some properties and applications of the same. The authors stated it as a new distribution thus different from the original unit Lindley (UL) distribution of [1].

A random variable $X \sim UL$ with parameter θ that is $UL(\theta)$ if its probability density function (pdf) and cumulative distribution function (cdf) are respectively given by

$$f_X(x \mid \theta) = \frac{\theta^2}{1+\theta} (1-x)^{-3} \exp\left(-\frac{\theta x}{1-x}\right), \qquad 0 < x < 1, \, \theta > 0.$$
(1)

and

$$F_X(x \mid \theta) = 1 - \left(1 - \frac{\theta x}{(1+\theta)(x-1)}\right) \exp\left(-\frac{\theta x}{1-x}\right), \qquad 0 < x < 1, \ \theta > 0.$$
(2)

On the other hand, if a random variable $Y \sim \text{NUL}$ (Mazuchelli et al., 2020) with parameter θ that is $\text{NUL}(\theta)$ then its pdf and cdf are respectively given as

$$f_Y(y \mid \theta) = \frac{\theta^2}{y^3(1+\theta)} \exp\left(-\frac{\theta \left(1-y\right)}{y}\right), \qquad 0 < y < 1, \, \theta > 0. \tag{3}$$

CONTACT Author. Email: subrata_stats@dibru.ac.in

and

$$F_Y(y \mid \theta) = \frac{\theta + y}{y(1+\theta)} \exp\left(-\frac{\theta \left(1-y\right)}{y}\right), \qquad 0 < y < 1, \ \theta > 0.$$
(4)

In the next section we provide main results which show that the NUL distribution is in fact complimentary of UL distribution.

2. Main Results

Theorem 2.1. If $X \sim UL(\theta)$, then $Y = 1 - X \sim NUL(\theta)$.

Proof:

$$F_Y(y \mid \theta) = F_{1-X}(y \mid \theta)$$

= $P(1 - X \le y)$
= $P(X > 1 - y)$
= $1 - P(X \le 1 - y)$
= $1 - \left[1 - \left(1 - \frac{\theta(1 - y)}{(1 + \theta)(-y)}\right) \exp\left(-\frac{\theta(1 - y)}{y}\right)\right]$
= $\frac{\theta + y}{y(1 + \theta)} \exp\left(-\frac{\theta(1 - y)}{y}\right).$

This is the cdf of $\text{NUL}(\theta)$ in equation (4). Hence proved.

As a result of this simple relation between the $UL(\theta)$ and $NUL(\theta)$ all properties of the later can be derived directly from the corresponding properties of the former. In what follows we show derivation of some of the important properties of $NUL(\theta)$ from those of $UL(\theta)$.

Theorem 2.2. The NUL(θ) is unimodal with mode at $Y = \frac{\theta}{3}$ for $\theta < 3$ and at Y = 1 for $\theta \geq 3$

Proof: It is known from [1] that $X \sim UL(\theta)$ is unimodal with mode at $X = 1 - \frac{\theta}{3}$ for $\theta < 3$ and at X = 0 for $\theta \ge 3$, Required result follows immediately from the fact that Y = 1 - X and $0 < X \le 1$.

Theorem 2.3. $NUL(\theta)$ belongs to exponential family.

Proof: It is known from [1] that the pdf of $X \sim UL(\theta)$ in equation (1) belongs to exponential family

$$f_X(x \mid \theta) = \exp\left[Q(\theta) T_X(x) + D(\theta) + S_X(x)\right].$$

with
$$Q(\theta) = -\theta, T_X(x) = x/(1-x), D(\theta) = \log\left[\frac{\theta^2}{1+\theta}\right], S_X(x) = \log(1-x)^{-3}.$$

Now using the transformation Y = 1 - X and $0 < X \le 1$, it is easy to verify that $\text{NUL}(\theta)$ with pdf in equation (3) can be rewritten as

$$f_Y(y \mid \theta) = \exp\left[Q(\theta) T_Y(y) + D(\theta) + S_Y(y)\right],$$

where $Q(\theta) = -\theta, T_Y(y) = T_X(1-y) = \frac{1-y}{y}, D(\theta) = \log\left[\frac{\theta^2}{1+\theta}\right], S_Y(y) = S_X(1-y) = \{\log(y)\}^{-3}$. Hence, belongs to the exponential family.

Theorem 2.4. The CDF of the $NUL(\theta)$ is convex for $\theta > 3$.

Proof: We know that $UL(\theta)$ is concave for $\theta > 3$ ([1]). Now

$$\frac{\mathrm{d}^2}{\mathrm{d}y^2} F_Y(y \mid \theta) = \frac{\mathrm{d}}{\mathrm{d}y} f_Y(y \mid \theta)$$
$$= \frac{\mathrm{d}}{\mathrm{d}y} f_X(1 - y \mid \theta)$$
$$= -f_X'(1 - y \mid \theta)$$
$$= -\frac{\mathrm{d}^2}{\mathrm{d}y^2} F_X(1 - y \mid \theta)$$

But because of the concavity of $UL(\theta)$ for $\theta > 3$ we have

$$\frac{d^2}{\mathrm{dx}^2}F_X(x\mid\theta) < 0.$$

Thus for $\theta > 3 \frac{d^2}{dy^2} F_Y(y \mid \theta) > 0$. Hence proved.

Theorem 2.5. The mean of the $NUL(\theta)$ is $\frac{\theta}{\theta+1}$.

Proof: We know that for $UL(\theta)$ the mean is $\frac{1}{\theta+1}$ ([1]). Hence mean of $Y \sim NUL(\theta)$ can be obtained as

$$E(Y) = E(1 - X) = 1 - \frac{1}{\theta + 1} = \frac{\theta}{\theta + 1}$$

In fact the k^{th} moment for $Y \sim \text{NUL}(\theta)$ can be written by using $E(Y^k) = E(1-X)^k$.

3. Conclusion

In this short article it is shown that the recently reported new unit Lindley distribution ([2]) is related to the original Unit Lindley distribution ([1]). All properties derived for either one can be easily translated for that of the other. Moreover the regression modeling of the mean of Y = 1 - X for $X \sim UL(\theta)$ distribution is equivalent to the regression modelling mean of $Y \sim NUL(\theta)$ distribution. As such the new unit Lindley distribution. It should be noted here that for any continuous random variable X with distribution in (0, 1), Y = 1 - X too will have a distribution in same range.

Acknowledgement(s)

The author acknowledges the editor and the reviewers for their suggestions.

4. Reference

References

- Mazucheli, J., Menezes, A.F.B., and Chakraborty, S.(2019). On the one parameter unit-Lindley distribution and its associated regression model for proportion data. Journal of Applied Statistics, 46(4), 700-714
- [2] Mazucheli, J., Bapat, S.R., Menezes, A.F.B. (2020). A new one-parameter unit-Lindley distribution. Chilean Journal of Statistics, 11(1), 53-67.